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COMMERCIAL BANKS AND THEIR
BUSINESS LOAN PORTFOLIOS: THIS RECOVERY
AND THE FUTURE

by

Sydney Smith Hicks

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Research Paper

FEDERAL RESERVE BANK OF DALLAS

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I. Introduction

According to many observers this recovery period has been
very different from previous recoveries in the sense that business
loans at commercial banks have recovered so slowly. Economists worry
about the strength of business loans because they indicate the strength
of business spending. The stronger business spending, the stronger
will be our recovery. Generally, bankers were mystified by the lack of
strength in business loans in 1975 and the continued recovery weakness
of business loans at the large banks. No doubt banks over these last
few years have maintained relatively more liquid portfolios than they
otherwise would have done, due to this expectation of a rebound in
business borrowing.

The existing models of business borrowing generally did not
predict the decline in business loans in 1975. A better forecast of
business borrowing would have enabled bankers to improve profitability
by enabling them to make more accurate portfolio decisions. The model
presented here and estimated through 1974, forecasts the decline in
1975 and tracks the current period fairly well.

In order to assess the cause of this weakness in business
borrowing a simple demand and supply model for the business loan market

is examined after prior studies of business loan behavior are
discussed in Section II. Based upon this model, equations for the
change in business borrowing at large commercial banks and for the
change in business borrowing from small commercial banks are esti-
mated and discussed in Section IV through VI. Major results of this
study are that (1) large bank and small bank markets are structurally
dissimilar, and (2) superior forecasts for total business loans can
be achieved by forecasting from large and small bank equations.

II. Previous Studies of Business Loan Behavior

Prior studies of business loan behavior generally fall into
two categories: demand studies or demand and supply studies. Four
recent studies--those by Harris [5], Goldfeld [4], Hendershott [6], and
the FMP model [10]--are summarized on Table 1. Many studies of busi-
ness loan behavior mention very little, if any, theoretical justification
for the inclusion of certain explanatory variables. Thus, what is
notable about the four lists of demand explanatory variables is their
diversity. When there is general agreement that the variable should be
included, there is lack of agreement about whether or not the variable
should enter in level or first difference form. The FMP model includes
the level of inventories; Hendershott and Harris include them in first
difference form. When the lagged business loan variable is included,
it is in level form in Goldfeld's study, in first difference form in
Hendershott's study, but is in combination with another variable in the
FMP formulation.

In the case of the interest rate variable in the demand
specifications, the disagreement is more complex. First, there is relatively

Table 1

Previous Studies' Explanatory Variables for
the Change in Business Loans

Demand Studies

Harris (1976)

ΔBook value of business inventories *
ΔBusiness fixed investment
Δ(Prime rate—commercial paper rate)
ΔCash flow

Goldfeld (1969)

Business loans lagged one period
Prime rate
Treasury bill rate
Quarterly dividend payments
Business sales
Time deposits lagged one period

Demand and Supply Studies

Hendershott Model (1968)

Demand
ΔBook value of business inventories
ΔCommercial loan rate
ΔBusiness loans lagged one period

Supply (variables determining Δ commercial loan rate)
Corporate Aaa rate
Monetary base
ΔBusiness loans lagged one period
Commercial loan rate lagged one period

FMP Model (1969)

Demand
Business inventories
Inventory adjustment factor
Expenditures on producers' durables
Expenditures on non-residential structures
GNP minus total investment (current and lagged)
(Treasury bill rate — commercial loan rate) (ΔTotal business product)
(Corporate Aaa rate — commercial loan rate) (ΔTotal business product)
(Amount of total investment adjusted for the inventory valuation
        adjustment — ΔBusiness loans), lagged one period

Supply (variables determining commercial loan rate)
Commercial and industrial loans/demand plus time deposits
Corporate bond rate
ΔFederal Reserve discount rate
Commercial paper rate, current and lagged one through five periods

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*   The symbol Δ stands for "change in".

little agreement regarding which rate or rates should be included. The
second issue is whether the chosen rate should enter in level form, first
difference form, or in deviation from another rate. And thirdly, one
of the models converts the interest rate variable into dollar terms,
whereas the other models use percentage terms.

The supply specifications contained in the Hendershott and
FMP models also display diversity. The only variable upon which both
models agree is that the corporate long-term bond rate should be in-
cluded. It is the only interest rate in Hendershott's formulation,
whereas the FMP model includes three different interest rates. The
quantity constraint variable is the monetary base in Hendershott's
model and is the ratio of business loans to the sum of demand and time
deposits in the FMP model.

In 1976, Harris reestimated Goldfeld and Hendershott's models.
With these reestimations of Goldfeld and Hendershott, his own model, and
simulations of the FMP model, Harris generated forecasts of 1975 business
loan behavior. The Goldfeld, Hendershott, and FMP models underpredicted
the 1975 decline by $24 billion, $7 billion, and $8 billion with root-
mean-square-errors of 6.56, 2.35, and 2.41, respectively.1/ Harris over-
predicted the decline by $.5 billion with a root-mean-square-error of
1.09. With the data base used in this study, the Harris model was re-
estimated and forecasts for 1975 were generated. The reestimated Harris

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1/ The root-mean-square-error criterion was used to judge the superiority
    of the forecasts and is defined as follows:

$$r\text{-}m\text{-}s\text{-}e = \sqrt{\left[\left(\sum_{i=1}^{4} e_i^2\right) / 4\right]}$$

where e_i is the error, or actual less predicted, in each period.

model still overpredicts the decline in 1975, but by an increased
average error of \$3.4 billion and a root-mean-square-error of 1.64.

Only the Harris model captured the extraordinary loan weak-
ness in 1975. Harris' major conclusion about this period was that
business loans were weak because of the lack in strength of inventory
spending and because there was an exceptional recovery in business cash
flows. Inventory spending and cash flows are demand variables. Supply
variables played no role in Harris' model and thus were not causative factors.

III. The Business Loan Market

In order to understand the business loan market, an exam-
ination of the portfolios of the participants is necessary. Although
commercial banks and nonfinancial businesses have very complex balance
sheets, only simple representations are used as the basis for this study.
Table 2 contains a concise summary of the model as well as an abbreviation
key.

Nonfinancial business firms can be characterized as financing
positions in cash (CSH), inventories (INV), and/or fixed capital (CAP) by
means of loans from commercial banks (BL), other liabilities which can be
short or long term (OL) and net worth (NW). The balance sheet constraint
for these firms is (Table 2, Equation 4)

$$\overline{NW}_F = CSH_F + \overline{INV} + \overline{CAP} - BL - OL.$$

Assume that at a given point in time the amounts of fixed capital (CAP) and
inventories (INV) the firm has are known to it, as well as the volume of
retained earnings or net worth (NW_F). Given these three quantities, the
level of bank loans demanded by the business firms (BL^d) depends upon the
interest rate charged by banks (the prime rate, r_p) and the interest rates

Table 2

A Simple Model of Business Loan Determination

<u>Model</u>

(1) $BL^d = d(r_p, r_{cp}, r_{Aaa}, INV, CAP, NW_F, BL_{t-1})$; $d_1, d_6 < 0$
$$d_2, d_3, d_4, d_5, d_7 > 0$$

(2) $BL^s = s(r_p, r_T, RAM, TLI, BL_{t-1})$; $s_2, s_3 < 0$
$$s_1, s_4, s_5 > 0$$

(3) $BL^d = BL^s = BL$

(4) $NW_F = CSH_F + INV + CAP - BL$

(5) $NW_B = CSH_B + R + L + I - DL$

(6) $\Delta CAP = BFI$

(7) $\Delta NW_F = CF$

<u>Model Solution for BL</u>

Level Form:

(8) $BL = f(r_{cp}, r_{Aaa}, r_T, RAM, TLI, INV, CAP, NW, BL_{t-1})$

First Difference Form:

(9) $\Delta BL = g(\Delta r_{cp}, \Delta r_{Aaa}, \Delta r_T, \Delta RAM, \Delta TLI, \Delta INV, BFI, CF, \Delta BL_{t-1}, Constant)$

<u>Model Key</u>

BFI	Business fixed investment
CAP	Capital
CF	Cash flow
CSH_B	Commercial bank cash
CSH_F	Nonfinancial business cash
DL	Deposit liabilities
I	Securities
INV	Inventory investments
L	Total Loans
NW_B	Commercial bank net worth
NW_F	Nonfinancial business net worth
OL	Other Liabilities
RAM	Reserve Adjustment Magnitude
r_{Aaa}	Corporate Aaa rate
r_p	Prime rate
r_t	Treasury 3- to 6- month bill rate
TLI	Total loans and investments , L + I

on other short-and long-term liabilities firms can issue (the

commercial paper rate, r_{cp}, and the long-term bond rate, r_{Aaa}).2/

The quantity of business loans demanded from banks varies

inversely with the prime rate. However, the demand for business loans

varies positively with interest rates on other types of liabilities, the

level of business inventories and the level of fixed capital. It varies

negatively with net worth.

It is possible that business firms do not adjust their bank loans

completely to equilibrium values within one period. This partial adjust-

ment may be the result of incomplete information and transactions costs.

As a result, some portion of the volume of loans desired but not under-

taken in the current period will be undertaken in the next period. Con-

sequently, the past level of loans $[BL_{t-1}]$ positively affects the demand

function for loans in the current period. Another reason for including this

variable is the bank-customer relationship. Business firms may borrow more

today, other factors being equal, in order to assure themselves of future

loan availability. Consequently, current loan demand depends on expected

future loan levels. Furthermore, if future loan levels are a function of

the past loan level, then BL_{t-1} is an explanatory variable in the demand

equation.3/

2/ The level of cash holds by the firms is determined as a residual once
the other factors on the balance sheet are known.

3/ The bank-customer relationship was introduced to the literature
by Donald R. Hodgman [9]. It has been extensively discussed and
tested by J. H. Wood [12].

Turning to the banking sector, banks can be characterized as financing positions in cash (CSH_B), reserves, (R), loans (L), and securities (I) by means of deposit liabilities (DL) and net worth (NW_B). The balance sheet constraint for the commercial banks is (Table 2, Equation 5)

$$\overline{NW_B} = CSH_B + \overline{R} + L + I - \overline{DL}.$$

In the current time period, bankers decide how many deposit liabilities and then set interest rates on those deposits to attract the funds. After subtracting required reserves (R) from the deposit liabilities (DL) and adding to that result the current amount of net worth (NW_B), the banks are assumed to allocate their "disposable assets" between securities (I) and loans (L) based on alternative rates of return on each.[4]   If the volume of excess reserves is small for the whole system, then the "disposable assets" (or the portfolio constraint variable) can be measured either as the sum of deposit liabilities plus net worth less reserves or as total loans and securities. The second approach is followed here, and thus the abbreviation for the port-folio constraint variable is TLI. An increase in the size of this port-folio constraint variable will increase holdings of both loans and securities.

Given the portfolio constraint variable, the amount of funds allocated by commercial banks to business loans is de-termined by what the banks can charge on the loans (the prime rate,

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[4]   "Disposable Assets" is a term used by William C. Brainard and James Tobin [2]. Brainard and Tobin make allowances for possible differ-ences in the effect of time deposits and demand deposits on loan supply in their theoretical model. This complication is ignored here. The volume of cash is determined once all the other magni-tudes are known; thus, the balance sheet constraint is satisfied.

r_p) and what the banks could earn on security investments (represented by the Treasury bill interest rate, r_T). When the prime rate increases and other factors remain the same, banks will increase the quantity of business loans supplied. When the Treasury bill rate increases, banks will decrease the supply of business loans because of the more attractive return on alternative investments.

The banks' allocation of total earning assets between business loans and other investments also depends on bank liquidity, which is affected to some extent by reserve requirements. For example, a bank facing a 5-percent reserve requirement would hold 5 cents in required reserves against $1 of deposits; if the $1 deposit was withdrawn, the bank would have to liquidate 95 cents of earning assets. A bank with a 15-percent reserve requirement would hold 15 cents in required reserves and would need to liquidate only 85 cents of such assets. Thus, when reserve requirements are low, it behooves the banker to be invested more heavily in securities than loans because of the relative liquidity of securities. The higher the reserve requirement, the less need there is for liquidity and the greater loans should be relative to securities.

A variable used previously in studies of the money supply process to measure the effects on reserves of changes in required reserves is the reserve adjustment magnitude, or RAM.5/ The reserve adjustment

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5/ RAM is discussed in detail by Leonall C. Andersen and Jerry L. Jordan [1] and by Albert E. Burger and Robert H. Rasche [3]. RAM was originally calculated so that a comprehensive variable could be constructed to measure the total impact of Federal Reserve policy on the monetary aggregates. The monetary base, which includes RAM, would then reflect the extent of open market operation, borrowing at the discount window, and reserve requirement changes.

magnitude translates changes in reserve requirements relative to a base period into dollars of reserves freed up or absorbed. An increase in reserve requirements reduces RAM and, thus, should lead to an increase in business loans relative to securities because the total earning asset portfolio can be less liquid.6/

Finally, the lagged level of business loans (BL_{t-1}) may affect the current level of business loans supplied by banks. Banks may not instantaneously adjust to desired levels the business loans they supply. This may be the case if, for example, information is incomplete and there are transactions costs in adjusting. The presumption is made that some portion of any desired increase in the supply of business loans not accomplished today will be undertaken in the next period. Consequently, the relation between last period's loan levels and today's loan levels is positive.

The quantity of business loans at any point in time is such that the amount supplied equals the amount demanded. This quantity is obtained from the simultaneous solution of Equations (1) through (3), which yields Equation (6). Model Equation (6) cannot be estimated as it is because there are no accurate measures of the fixed capital stock (CAP) or the net worth (NW) of nonfinancial businesses. However, business fixed investment (BFI) measures the addition to capital stock each period, and an indication of the addition to net worth each

6/ In a simplified model, $RAM_t = (r_o - r_t) D_{t-2}$, where r_o is the required reserve ratio in the base period, r_t is the required reserve ratio in the current period, and D_{t-2} is the level of deposits two periods ago. Because the model for business loans is estimated in first-difference form, the first difference of RAM is used in the estimated model. The change in RAM captures the dollar amount of reserves freed or absorbed by concurrent changes in reserve requirements, adjusted for shifts in deposits among banks.

period is undistributed corporate profits (CF).7/ As a result, the
equation was estimated in first-difference form, as represented in
Equation (7).

To the extent that a bank responds to an increase in busi-
ness loan demand by selling more liabilities, a portion of the portfolio
constraint variable becomes endogenous. If this were true for all banks,
we could not be sure whether an increase in the aggregate portfolio con-
straint variable led to an increase in business loans or vice versa.
However, deposit liabilities and, thus, total earning assets for the whole
banking system are importantly constrained by the total amount of reserve
money supplied by the Federal Reserve System.8/ That the assumption of an
exogenous portfolio constraint variable (ΔTLI) is a reasonable assumption has
been confirmed by the two-stage least-squares estimates. The two-stage
estimates attribute at least as much importance to the supply effects of total
earning assets as do ordinary least-squares estimates.9/

IV. Estimation of the Model

The model's equation for the change in business loans was
estimated for all commercial banks, for large commercial banks (the

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7/ The CF variable is undistributed corporate profits plus the inventory
    valuation adjustment and depreciation. There exists the possibility of
    measurement error in the business loan series due to judgments regarding
    loan classification. Consequently, a constant should be and was added
    for econometric reasons. For a discussion of these problems, see Robert
    S. Pindyck and Daniel L. Rubinfeld [12, pp. 128-129].

8/ The problem of simultaneous-equation bias in the ordinary least-squares
    estimation used here would remain if the Federal Reserve tended to sup-
    ply or withdraw reserves automatically in response to variations in bank
    loan demand. Since the Trading Desk of the Federal Reserve follows an
    interest rate target between the monthly meetings of the Federal Open
    Market Committee, this could be a problem for data covering relatively
    short periods. But over the quarterly intervals used in this study there
    is often substantial movement in short-term interest rates, so total earn-
    ing assets of banks can still be considered exogenous.

9/ Hicks [8, pp. 15-16].

weekly reporting banks), and for small commercial banks (all banks
excluding the weekly reporting banks). [Table 3] The equations were
estimated from 1960III-1974IV; 1960III represents the beginning of the
period for which bank data disaggregated by size is available and 1974IV
is the last data point before the seemingly unusual business loan
behavior began. All of the regressions are significant (as measured
by the F statistic), and the Durbin-Watson statistics (D-W) are close to
2.0 indicating very little residual autocorrelation.

Before discussing the estimated coefficients it is worth-
while to examine whether or not disaggregation of the business loan
equations is appropriate. To find out, another regression was estimated
based on a test developed by Zellner.10/ In general functional form the
change in business loans at all banks can be explained as follows:

$$\Delta BLA = f(C, \Delta r_{cp}, \Delta r_{Aaa}, \Delta r_T, \Delta RAM, \Delta TLIA, \Delta INV,$$
$$BFI, CF, \Delta BLA_{t-1}, \Delta TLIL, \Delta BLL_{t-1}).$$

Two variables from the large bank equation, $\Delta TLIL$, and ΔBLL_{t-1}, were
added to the aggregate model equation of Table 3. If the estimated
coefficients on these variables are significantly different from zero,
then disaggregation is appropriate. The equation was estimated over the
period 1960III-1974IV. The coefficient of $\Delta TLIL$ was positive with a
t-statistic of 1.14, and the coefficient of ΔBLL_{t-1} was positive with
a t-statistic of 4.24. The R^2 was .9216 (with an adjusted R^2 of .9029).
An F-test conducted on the hypothesis that both coefficients equaled
zero resulted in the rejection of the hypothesis.

To test the stability of this result, the equation was
estimated over sample periods extended by one year at a time. The

10/ See Zellner [13].

Table 3

BUSINESS LOAN EQUATION

Explanatory Variable	Estimates		
	All Banks	Large Banks	Small Banks
Constant (C)	.128 (.21)	.097 (.19)	-.381 (-2.13)
Change in:			
Commercial paper rate (Δr_{cp})	.932 (2.51)	1.232 (4.00)	-.063 (-.59)
Long-term corporate bond rate (Δr_{Aaa})	.003 (.003)	-.495 (-.69)	-.131 (-.54)
Reserve adjustment magnitude (ΔRAM)	.588 (2.04)	.717 (2.80)	-.112 (-1.24)
Treasury bill rate (Δr_{T})	-.141 (-.30)	-.311 (-.74)	.151 (.95)
Total loans and investments at all banks $(\Delta TLIA)$	.182 (4.79)	--	--
Total loans and investments at large banks $(\Delta TLIL)$	--	.178 (4.52)	--
Total loans and investments at small banks $(\Delta TLIS)$	--	--	.065 (3.85)
Inventories (ΔINV)	.438 (3.85)	.339 (3.29)	.053 (1.74)
Business fixed investment (BFI)	-.006 (-.56)	-.001 (-.13)	.011 (3.08)
Corporate cash flow (CF)*	.002 (.08)	-.003 (-.15)	-.008 (-1.05)
Lagged change in:			
Business loans at all banks (ΔBLA_{t-1})	.215 (2.19)	--	--
Business loans at large banks (ΔBLL_{t-1})	--	.228 (2.58)	--
Business loans at small banks (ΔBLS_{t-1})	--	--	-.162 (-1.06)
R^2 (measure of adequacy of fit)	.887	.862	.841
R^2 Adjusted	.866	.836	.811
D-W (Durbin-Watson autocorrelation test statistic)	1.999	1.834	1.982
SE (standard of error of the regression)	.843	.758	.258

* Undistributed corporate profits plus the inventory valuation adjustment and depreciation.

NOTE: Equations estimated for 1960III through 1974IV
Figures in parentheses are t-statistics of the regression coefficients.

significance of ΔBLL_{t-1} fell while the significance of $\Delta TLIL$ grew. In summary, disaggregation of the aggregate business loan market yields more information than the aggregate equation for business loan behavior.

V. A Comparison of Large and Small Bank Business Loan Markets

A few interesting differences and similarities between small and large bank business loan markets can be noted by comparing the coefficient estimates in Table 3. Generally, the coefficient estimates have the positive or negative signs economists would expect, given the prior behavioral assumptions. Despite a great degree of collinearity among the variables (which reduces t-statistics), many of the explanatory variables are still significant.

At the large banks, for example, an increase in the commercial paper rate (Δr_{cp}) of one percentage point will increase business loans by $1.232 billion (as the alternative means of financing becomes more expensive). If the reserve requirements are lowered releasing $1 billion in reserves, business loans increase $.72 billion. If either inventories (ΔINV) or total loans and investments ($\Delta TLIL$) increase by $1 billion, business loans increase by $.34 and $.18 billion, respectively. Despite the view of some large banks that their business loans increase when business fixed investment (BFI) rises, these results do not indicate this this.[11] Contrary to Harris' study, these results do not indicate a significant impact of cash flows on the change in business loans. And finally, the combined lagged adjustment of large banks and their customers results in

11/ See Herman [7].

a significantly positive lagged effect of last period's change
in business loans on this period's change in business loans. A $1
billion increase in business loans last period will increase this
period's business loans by about one quarter of a billion dollars.

The results for small banks are similar with respect to
sign, but the magnitudes of the coefficients are much different. Mul-
ticollinearity among the interest rates (two short-term and one long-
term) appears to be more of a problem in the small bank equation esti-
mates; none of the interest rates have coefficients significantly dif-
ferent from zero. As with large banks, a billion dollar increase in
either inventories or total loans and investments increases business
loans significantly by $.07 billion or $.05 billion, respectively.
A rather surprising result is the negative coefficient on RAM; in other
words, when reserve requirements are lowered thus releasing reserves and
increasing RAM, the change in business loans is reduced. There is some
comfort in the fact that the coefficient is not significant at the 95-
percent level. Unlike the large bank regression, (1) rising levels of
business fixed investment add to the current change in business loans,
(2) increasing cash flows measurably depress borrowing at small banks
(the t level is more negative, but still not significant) and (3) the
lagged effect of the past period's change in business loans does <u>not</u>
significantly affect the current change in business loans.

The lagged change in business loans coefficient can be
interpreted as a measure of the importance of the loan-customer rela-
tionship in the small and large bank markets. The insignificant
coefficient on ΔBLS_{t-1} does make sense if it is true that in the small
bank markets firms do not have much choice as to where to bank, and

the banks do not have much competition. In the large bank market,
there may be relatively more competition among bankers and more of the
large bank customers may have alternative financing options; as a result,
the loan-customer relationship may become more significant as a tool
for maintaining the banks' market shares.12/

VI. Forecasting Business Loans

Besides the fact that multicollinearity among many of the
explanatory variables did mean that some coefficients were unexpect-
edly insignificant (for example, the cash flow variable), collinearity
causes the estimates of the coefficients to change dramatically when
sample periods are updated and when data is revised. One or more
variables could be eliminated to reduce collinearity; this would also
reduce the number of variables which would have to be forecast before
a business loan prediction could be generated. However, eliminating
variables does result in specification error. To warrant confidence
in the coefficient estimates and the predictions based on them, any
specification that omits variables should predict outside the sample
period at least as well as the whole model. Otherwise, the specifica-
tion error introduced would be too costly for the gain in coefficient
stability.

In fact, some of the specifications of the model that
omitted some interest rate and/or RAM variables did predict 1975 better
than the whole model estimated through 1974IV, as judged by the root-

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12/ The insignificant coefficient on the lagged change in small bank
    business loans remained robust for sample period endpoints ranging
    from 1970-1977I. When the 1977II-1978I data was added, the co-
    efficient became significant and positive. Because the small bank
    equation is unstable in this period, more data is needed before this
    new result can be viewed as accurate.

mean-square-error statistic. From the alternative specifications of
the model estimated with data available in March 1978, the following
specifications for large and small banks (which subsequently will be
called the restricted model equations) minimized the r-m-s-e for 1975:

(a) $\Delta BLL = f' (C, \Delta r_{Aaa}, \Delta r_T, \Delta TLIL, \Delta INV, BFI, CF, \Delta BLL_{t-1})$

and

(b) $\Delta BLS = g' (C, \Delta r_{Aaa}, \Delta r_T, \Delta RAM, \Delta TLIS, \Delta INV, BFI, CF, \Delta BLS_{t-1})$. [12]/
The equation for large banks (a) excludes ΔRAM and Δr_{cp} from the theo-
retical model; only Δr_{cp} is removed from the small bank equation (b).
Tables 4 and 5 contain the root-mean-square-errors of various predic-
tion periods for the theoretical model and the restricted model, re-
spectively.

No matter what specification was examined, a superior total
forecast for 1975 was always made by forecasting the small and large
bank components and then adding them together. For example, this result
may be observed from the first line of Table 4. When the theoretical
model was estimated over estimation periods ending later than 1974IV,
the predictions for total business loans made from the disaggregated
small and large bank equations were generally better than aggregate
predictions. The aggregate predictions were substantially better than
disaggregated predictions only during periods when the structural

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12/ The coefficient estimates are presented in Hicks [8]. The Δr_{Baa}
variable was used instead of Δr_{Aaa} in the small bank regressions
because Δr_{Baa} probably proxies the long-term borrowing costs of
small bank customers better than Δr_{Aaa}. However, because of the
statistical tests conducted on the model in this paper, it was
necessary to use Δr_{Aaa} instead of the Δr_{Baa} variable.

TABLE 4

Root-Mean-Square-Error Statistics For the
Theoretical Model over Alternative Prediction Periods

Estimation Period	Prediction Period	Aggregate All Banks	Large Banks	Small Banks	Disaggregate* --All Banks
1960III-1974IV	1975I-1975IV	2.563	1.975	.548	2.424
1960III-1974IV	1975I-1976IV	3.116	2.465	.648	3.012
1960III-1974IV	1975I-1978I	2.515	2.159	.966	2.518
1960III-1975IV	1976I-1978I	2.328	1.814	1.227	2.436
1960III-1976IV	1977I-1978I	2.176	1.546	1.528	2.439
1960III-1977IV	1978I-	1.954	1.794	.276	1.518

* The disaggregate all bank r.m.s.e. statistics are generated from the errors of the individual large and small bank equations.

TABLE 5

Root-Mean-Square-Error Statistics For the
Restricted Model over Alternative Prediction Periods

Estimation Period	Perdiction Period	Large Banks	Small Banks	Disaggregate* --All Banks
1960III-1974IV	1975I-1975IV	1.038	.556	1.480
1960III-1974IV	1975I-1976IV	1.740	.660	2.165
1960III-1974IV	1975I-1978I	1.465	.953	1.926
1960III-1975IV	1976I-1978I	1.629	1.218	2.243
1960III-1976IV	1977I-1978I	1.055	1.519	2.061
1960III-1977IV	1978I-	1.656	.274	1.382

* The disaggregate all banks r.m.s.e. statistics are generated from the errors of the individual large and small bank equations.

equations were unstable and thus could not be considered reliable.14/

There appear to be some sizable gains in prediction accuracy
for large bank business loans when the restricted model is used, re-
gardless of the period of estimation and forecast period. This does
not seem to be true for the small banks prediction errors; the r-m-s-e
statistics are very similar. Although the errors are generally lower
for the small bank predictions, they are 58 and 60 percent of the average
quarterly change in small bank business loans during 1975 for the
theoretical and restricted models, respectively. The large bank predic-
tions are 75 and 40 percent of the average change in large bank business
loans in 1975 for the theoretical and restricted models, respectively.

The total business loan root-mean-square-error statistics
improve dramatically from $2.56 billion for the aggregate bank theoretical
model in 1975, to $1.48 billion for the restricted disaggregate predictions.
The lowest r-m-s-e statistic for 1975 generated from prior studies was
1.64; the worst was 6.56. While the restricted model does a better job
predicting total business loans, the error does remain 89 percent of the
1975 average quarterly change in business loans.

The model does predict the decline in total business loans in
1975 better than prior models.15/ (Chart 1) Most of the weakness oc-
curred at the large banks, while changes in small bank business loans re-
mained stable. In 1976 and 1977 large bank predictions were good in the
sense of not missing consistently in the same direction. On the other
hand, the changes in small bank business loans in 1977 were consistently

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14/ See Hicks [8, p.16-17].

15/ The predictions were generated from the restricted model estimated
from 1960III-1974IV.

CHART 1. Actual Versus Predicted Changes in Business Loans
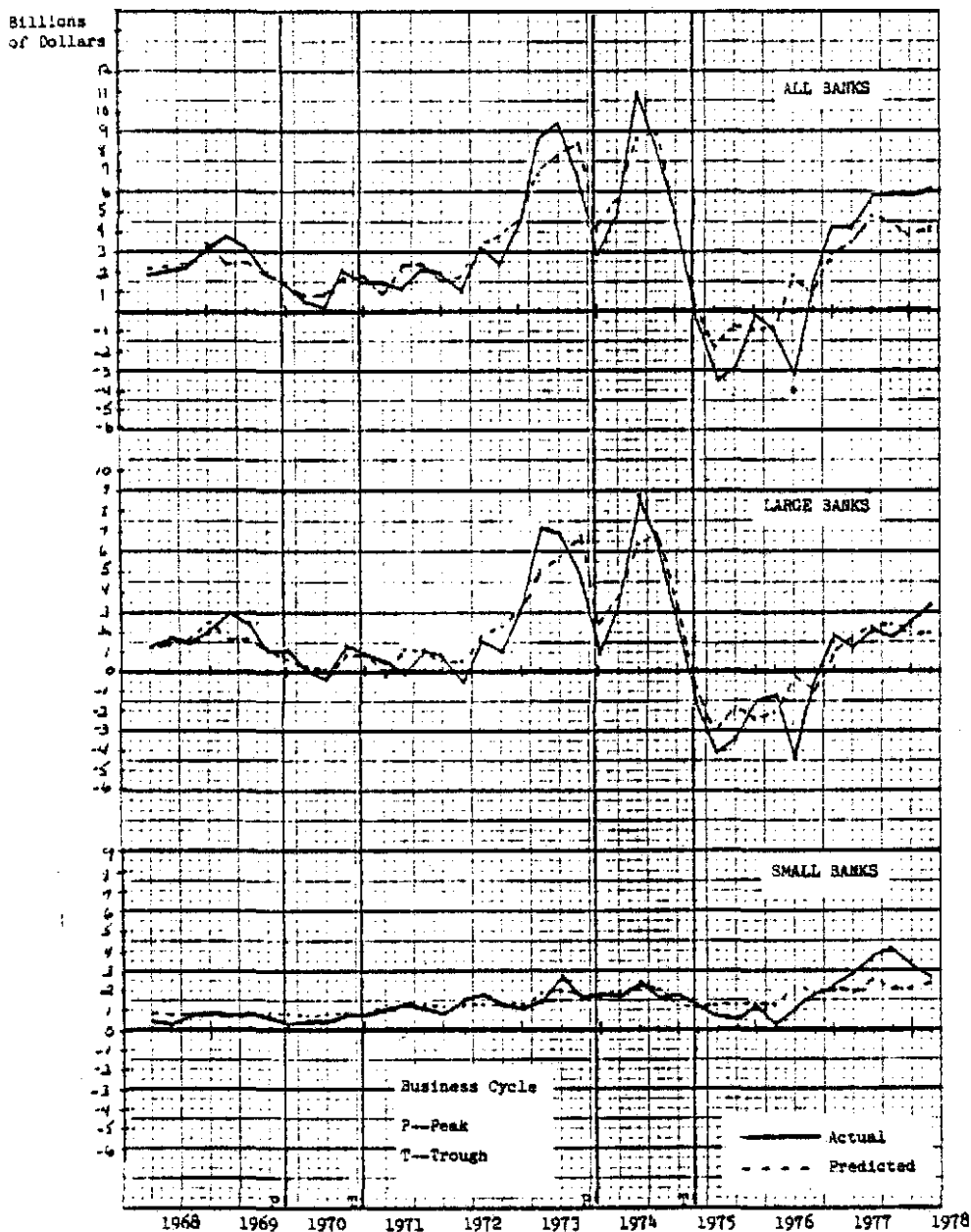
*Reclassification of loans as of March 31, 1976, lowered the change
in business loans by $1.2 billion in 1976-Q2.

NOTE: Predictions generated from model estimates for 1960-Q3
through 1974-Q4.

SOURCES: Federal Reserve Bank of St. Louis.
          Federal Reserve Bank of Dallas.

underestimated. Structural stability tests presented elsewhere
indicate relatively more structural instability in the current time
period for the small bank business loan market than for the large bank
market.16/

VII. <u>Conclusion</u>

In the case of the business loan market, aggregation of
small and large bank markets is not appropriate. Estimates of the
disaggregated large and small bank business loan equations provide
interesting similarities as well as dissimilarities. One of the most
interesting results is that last quarter's change in business loans
in the small bank market provides no significant information about to-
day's change in business loans, contrary to the results for the large bank
business loan market.

Understanding business loan behavior has proven to be a very
difficult task. Recognizing the structural diversity between large and
small markets rather dramatically increases the explanatory and predictive
power of the model. Modeling both the demand and supply sides of the
market yielded equations which, upon estimation, provided better pre-
dictions of 1975 than alternative formulations. Estimated through 1974,
the model predicts the decline in total business loans, as well as the
relative weakness in the large bank loan market. To the extent the model
increases the ability of bankers to predict business loan behavior,
bankers will be able to improve profitability by making more accurate port-
folio decisions.

16/ Hicks [8, p. 17-18].

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